

Nordic Summer School in Cancer Epidemiology

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Measures of disease frequency and effects

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Outline

Basic concepts

Frequency

Comparison

Age, period, etc.

Standardisation

Survival

Conclusion

Key references

- IS: dos Santos Silva, I. (1999).
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C&H: Clayton, D., Hills, M. (1993).
Statistical Models in Epidemiology. OUP, Oxford.

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- ▶ “study of the distribution and determinants of health related **states** and **events** in specified populations, ...” (Porta (ed.) Dictionary of Epidemiology, 2008)
- ▶ “discipline on principles of **occurrence** research in medicine” (Miettinen 1985)

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- ▶ **clinical** epidemiology – study of diagnosis, prognosis and effectiveness of therapies in patient populations
– basis of evidence-based medicine

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But what are **rate** and **occurrence**?

Cancer in Norden 1997 (NORDCAN)

Frequency of cancer (all sites excl. non-melanoma skin) in Nordic male populations expressed by different measures.

	New cases	Crude rate	ASR (World)	Cumul. risk	SIR
Denmark	11 787	452	281	27.8	104
Finland	10 058	<u>401</u>	269	26.5	101
Iceland	<u>633</u>	464	347	32.6	132
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- ▶ What do these measures mean?

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What are the **proportions** or/and **rates** of occurrence of these states and events?

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- ▶ What are the **excess** and **relative risks** for nulliparous compared to parous women?
- ▶ What is the **dose-response relationship** between occupational exposure to crystalline silica and the risk of getting lung cancer in terms of level and length of exposure?

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- ▶ **Average risks** of disease in large groups sharing common characteristics (like gender, age, smoking status) are estimable from appropriate epidemiologic studies by pertinent **measures of occurrence**.

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Challenge: *How to find a **comparable** group of unexposed?*

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Focus in this course: *observational*, and *longitudinal cohort* & *case-control* studies.

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 - Possibility of **confounding**: due to other determinants of the outcome, correlated with exposure.
 - * Challenges: **Valid**: and **efficient** non-randomized design and statistical analysis.

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Causal research \rightarrow longitudinal base preferred.

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 - ▶ Person-time calculation complicated.
 - ▶ Population-based annual (or 5-year period) incidence and mortality statistics:
 $Y \approx \text{mid-population} \times \text{length of period.}$

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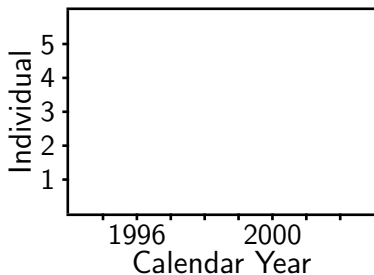
Also called **incidence density**.

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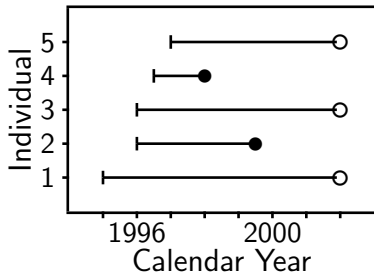
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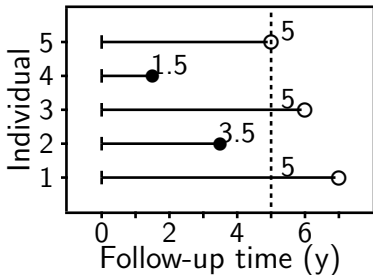
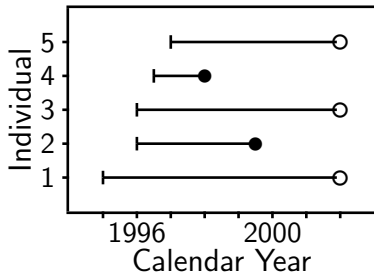
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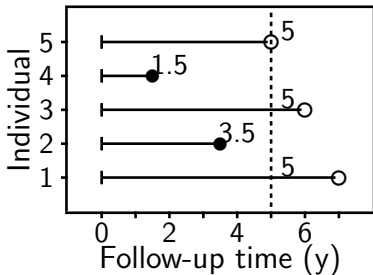
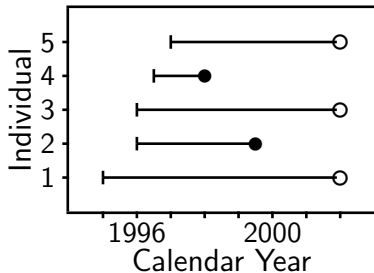
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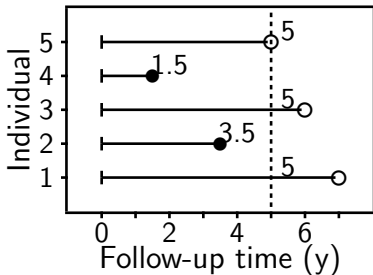
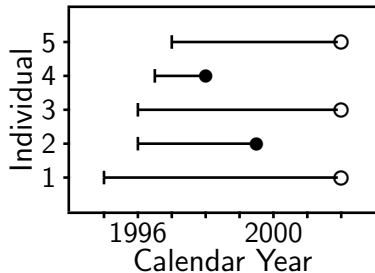
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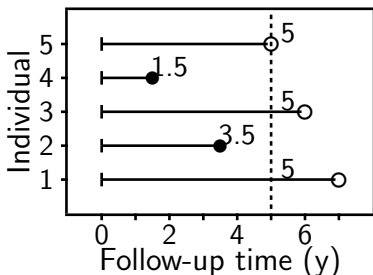
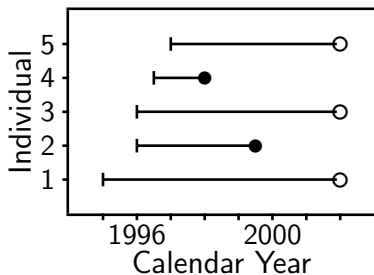


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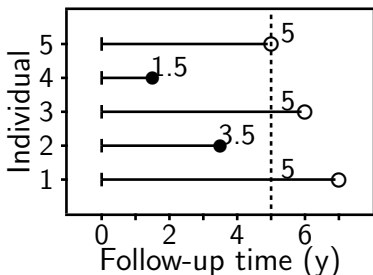
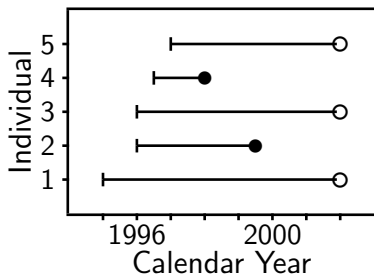
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In the constant hazard model with no competing risks:

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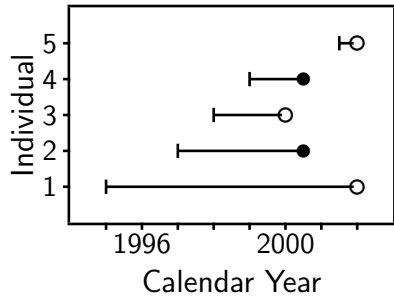
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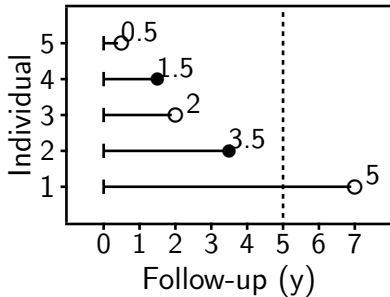
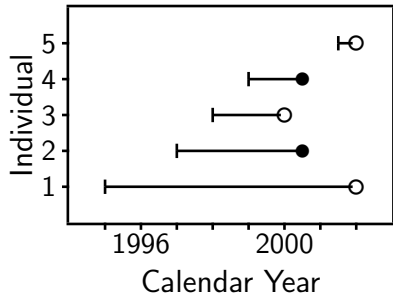
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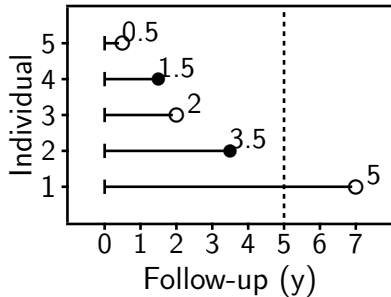
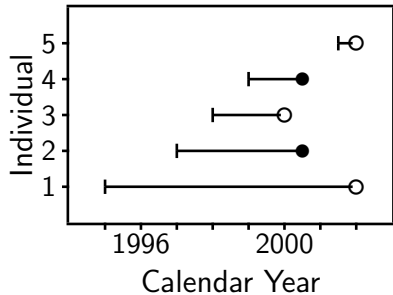
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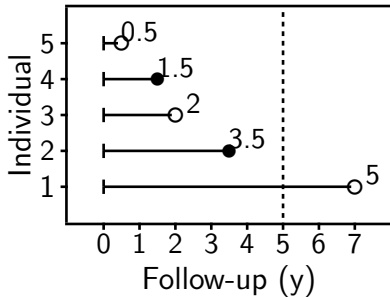
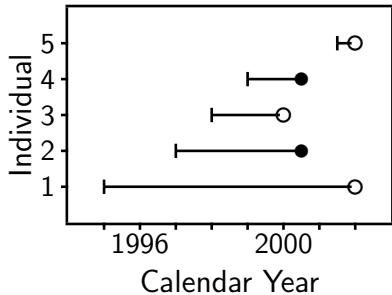


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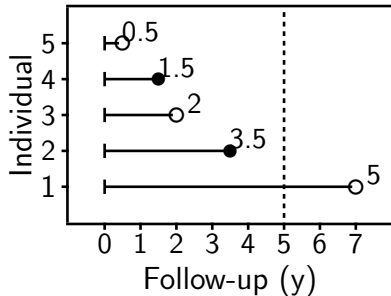
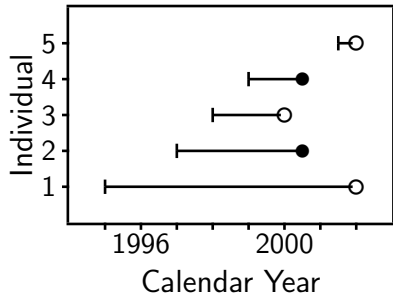
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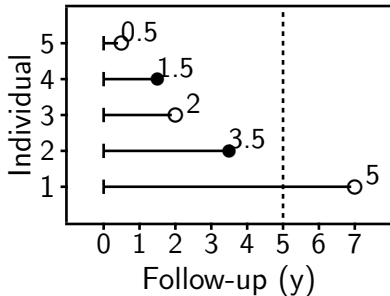
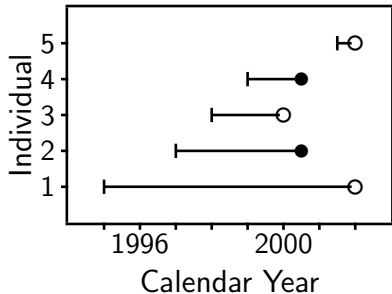


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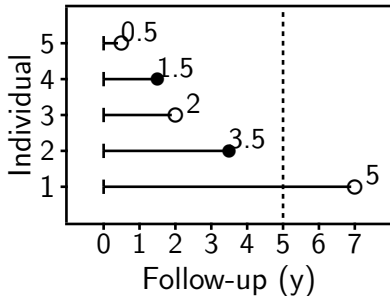
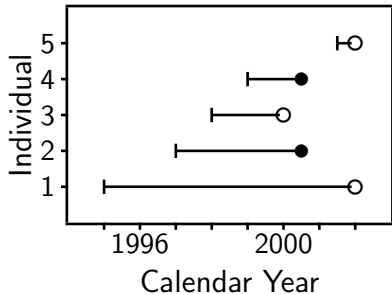
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NB. The actual study population often contains also some already affected, who thus do not belong to the population at risk. With rare outcomes the influence of this is small.

Male person-years in Finland 1991-95

Total male population (1000s) on 31 December by year:

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$$\begin{aligned}1992: \quad & \frac{1}{2} \times (2443 + 2457) \times 1 = 2450 \\ & \frac{1}{2} \times (2457 + 2482) \times 2 = 4937 \\ & \frac{1}{2} \times (2431 + 2492) \times 5 = 12307.5\end{aligned}$$

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$$\begin{aligned} 1992: & \quad \frac{1}{2} \times (2443 + 2457) \times 1 = 2450 \\ 1993-94: & \quad \frac{1}{2} \times (2457 + 2482) \times 2 = 4937 \\ & \quad \frac{1}{2} \times (2431 + 2492) \times 5 = 12307.5 \end{aligned}$$

Male person-years in Finland 1991-95

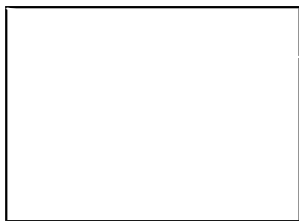
Total male population (1000s) on 31 December by year:

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Incidence proportion, rate, and odds (IS, Ex 4.5)



Time (t) \longrightarrow

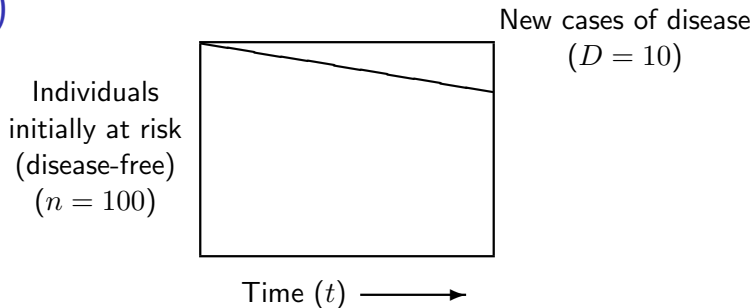
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Individuals
initially at risk
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($n = 100$)

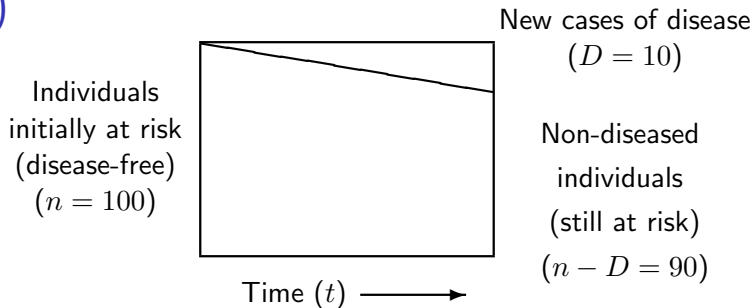


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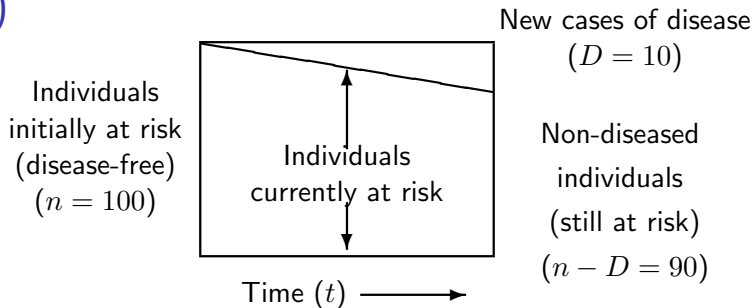
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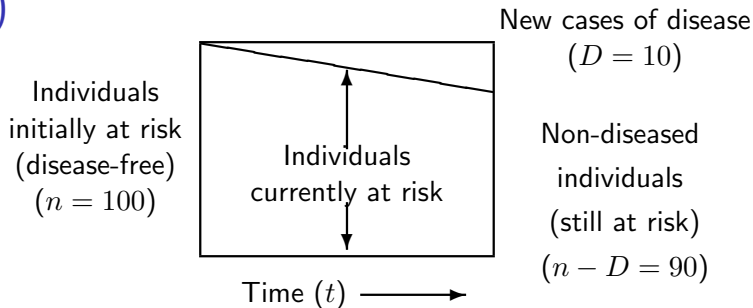
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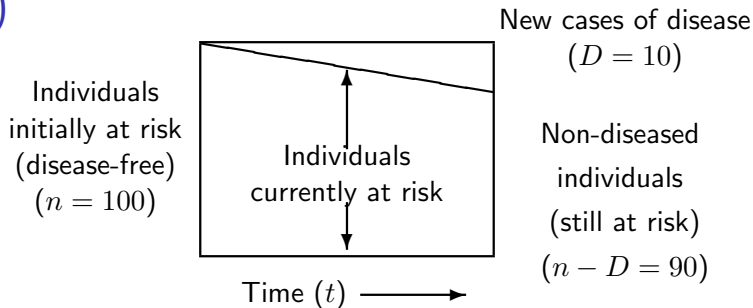


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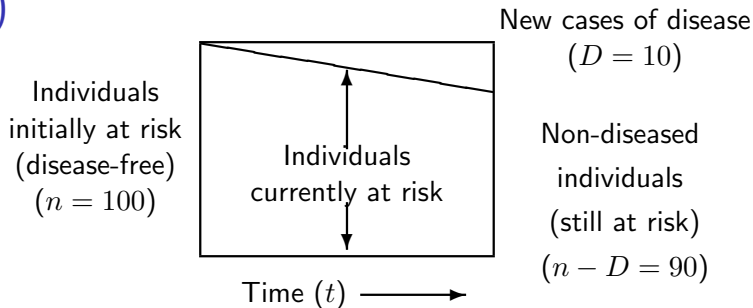
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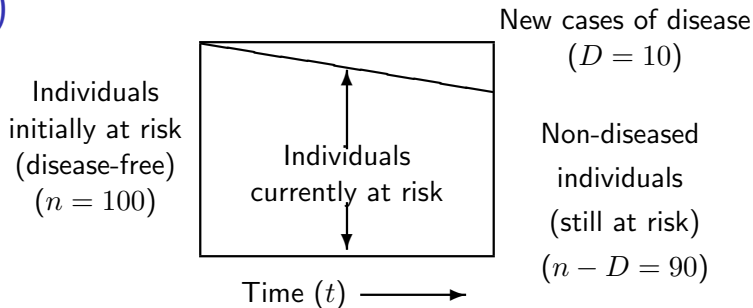
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The “**rare disease assumption**”.

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Mortality depends on the incidence and the **prognosis** or **case fatality** of the disease, *i.e.* the **survival** of those affected by it.

Mathematical concepts describing risks

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$\text{Exp}(\lambda)$ model \Rightarrow **Likelihood function** of λ is equivalent to that when number of cases D would be **Poisson**-distributed

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More about these issues in Bendix's lectures next week.

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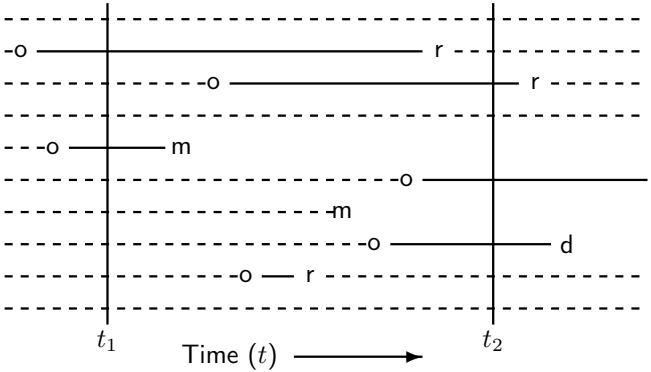
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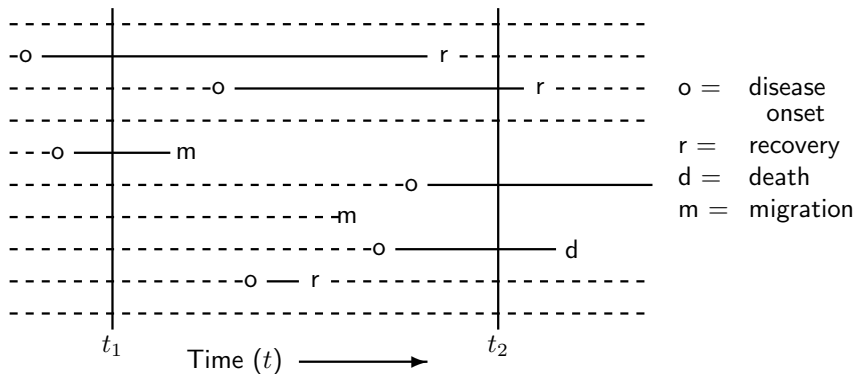
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Example 4.1 (IS: p. 59)



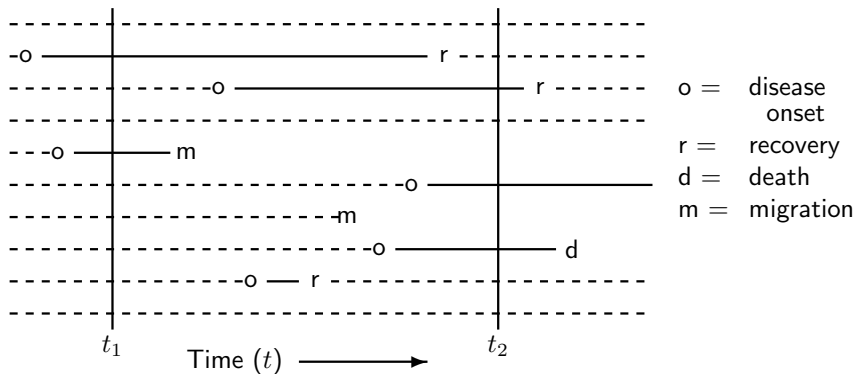
- o = disease onset
- r = recovery
- d = death
- m = migration

Example 4.1 (IS: p. 59)



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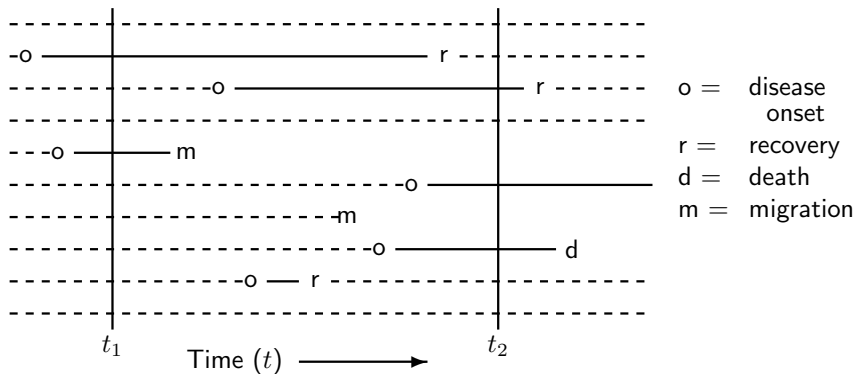
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Period prevalence: $5/8 = 0.62 = 62\%$

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The approximation works well, when $P < 0.1$ (10%).

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Often further classified by years since diagnosis.

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In causal studies these are used to estimate the **causal effect** of the exposure factor on the disease risk.

⇒ **comparative measures** \approx **effect measures**

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- ▶ Ratios – describe the **biological strength** of the exposure
- ▶ Differences – inform about its **public health importance**.

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Factor X has a stronger biological potency for disease A, but it has a greater public health importance for disease B.

Ratio measures in “rare diseases”

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$$\text{Inc. prop'n ratio} = \frac{30/4\,000}{60/16\,000} = \frac{7.5 \text{ per } 1\,000}{3.75 \text{ per } 1\,000} = 2.0000$$

$$\text{Inc. rate ratio} = \frac{30/7\,970 \text{ y}}{60/31\,940 \text{ y}} = \frac{3.76 \text{ per } 1\,000 \text{ y}}{1.88 \text{ per } 1\,000 \text{ y}} = 2.0038$$

$$\text{Inc. odds ratio} = \frac{30/(4\,000-30)}{60/(16\,000-60)} = \frac{0.00756}{0.00376} = 2.0076$$

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EF estimates the fraction out of all new cases among those exposed, which are “caused” by the exposure itself, and which thus could be “avoided” if the exposure were absent

Next time: Graphics of impact measures

Apply Bendix's R script on how to draw pictures to illustrate the concepts of excess fraction and population excess fraction with given RRs and prevalences of exposure.

Measures of potential impact (cont'd)

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Population EF and **population PF** combine these further with the *prevalence of exposure* in target population.

Effect of smoking on mortality by cause

(IS: Example 5.14, p. 98)

Underlying cause of death	Never smoked regularly Rate ^b	Current cigarette smoker Rate ^b	Rate ratio	Rate difference ^b	Excess fraction (%)
	(1)	(2)	(2)/(1)	(2) - (1)	$\frac{(2) - (1)}{(2)} \times 100$
Cancer					
All sites	305	656	2.2	351	54
Lung	14	209	14.9	195	93
Oesophagus	4	30	7.5	26	87
Bladder	13	30	2.3	17	57
Respiratory diseases (except cancer)	107	313	2.9	206	66
Vascular diseases	1037	1643	1.6	606	37
All causes	1706	3038	1.8	1332	44

^a Data from Doll *et al.*, 1994a.

^b Age-adjusted rates per 100 000 pyrs.

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Age is usually the strongest time-dependent determinant of health outcomes.

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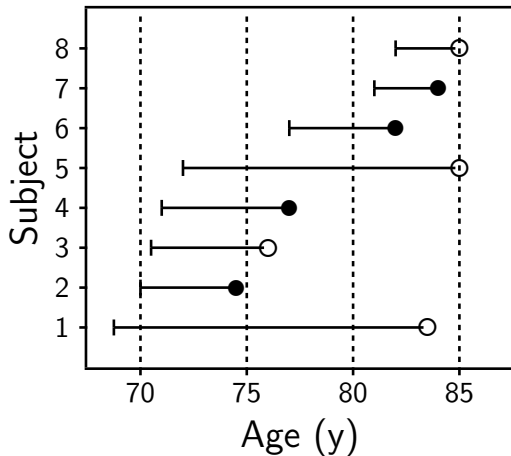
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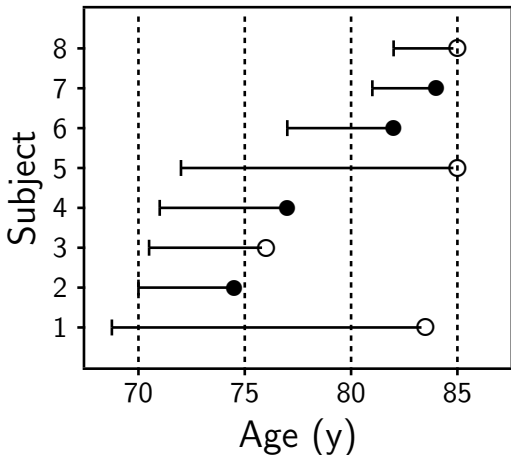
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Age is also often correlated with duration of “chronic” exposure (*e.g.* years of smoking).

Follow-up of a small geriatric cohort

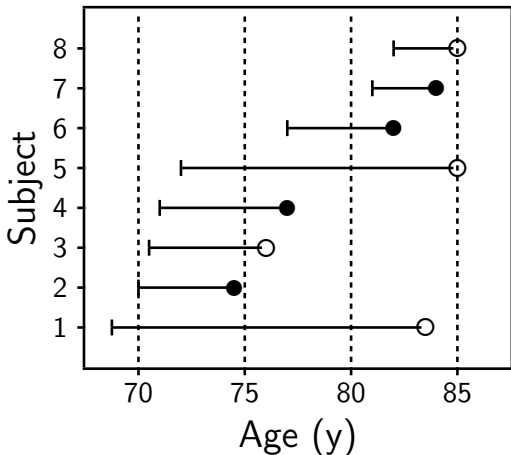


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Hides the fact that the “true” rate probably varies by age,
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$$I_k = \frac{\text{number of cases observed in ageband}}{\text{person-years contained in ageband}}$$

- ▶ Underlying assumption: **piecewise constant rates model**

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Subject	Ageband			Total
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	4.5	1.0	-	5.5
	4.0	2.0	-	6.0
5	3.0	5.0	5.0	13.0
6	-	3.0	2.0	5.0
7	-	-	3.0	3.0
8	-	-	3.0	3.0
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Calendar period	Age group (y)									
	40-44	45-49	50-54	55-59	60-64	65-69	70-74	75-79	80-84	85+
1953-57	21	61	119	209	276	340	295	279	193	93
1958-62	22	65	135	243	360	405	429	368	265	224
1963-67	24	61	143	258	395	487	509	479	430	280
1968-72	21	61	134	278	424	529	614	563	471	358
1973-77	16	50	134	251	413	541	629	580	490	392
1978-82	13	36	115	234	369	514	621	653	593	442
1983-87	11	31	74	186	347	450	566	635	592	447
1988-92	9	25	57	128	262	411	506	507	471	441
1993-97	7	22	48	106	188	329	467	533	487	367
1998-02	5	14	46	77	150	239	358	445	396	346

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Calendar period	Age group (y)									
	40-44	45-49	50-54	55-59	60-64	65-69	70-74	75-79	80-84	85+
1953-57	21	61	119	209	276	340	295	279	193	93
1958-62	22	65	135	243	360	405	429	368	265	224
1963-67	24	61	143	258	395	487	509	479	430	280
1968-72	21	61	134	278	424	529	614	563	471	358
1973-77	16	50	134	251	413	541	629	580	490	392
1978-82	13	36	115	234	369	514	621	653	593	442
1983-87	11	31	74	186	347	450	566	635	592	447
1988-92	9	25	57	128	262	411	506	507	471	441
1993-97	7	22	48	106	188	329	467	533	487	367
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- ▶ Rows: age-incidence pattern in different calendar periods.
- ▶ Columns: Trends of age-specific rates over calendar time.

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- ▶ Time trends inconsistent across age groups.

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- ▶ Analysis of rates by birth cohort reveals, how the level of incidence (or mortality) differs between successive generations – may reflect differences in risk factor levels.
- ▶ Often more informative about “true” age-incidence pattern than age-specific incidences of single calendar period.

Age-specific rates by birth cohort

Calendar period	Age group (y)								
	40-44	45-49	50-54	55-59	60-64	65-69	70-74	75-79	
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Adjustment for age and calendar time needed, e.g. by comparing *observed* to *expected* cases with SIR (see p. 76-79).

⇒ Cases and person-years in the study cohort must be split by more than one time scale (age).

Example (C&H, Tables 6.2 & 6.3, p. 54)

Entry and exit dates for a small cohort of four subjects

Subject	Born	Entry	Exit	Age at entry	Outcome
1	1904	1943	1952	39	Migrated
2	1924	1948	1955	24	Disease C
3	1914	1945	1961	31	Study ends
4	1920	1948	1956	28	Unrelated death

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Subject 1: Follow-up time spent in each ageband

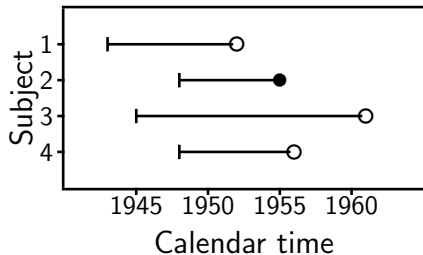
Age band	Date in	Date out	Time (years)
35–39	1943	1944	1
40–44	1944	1949	5
45–49	1949	1952	3

Example: (C&H, Figures 6.1 & 6.2, p. 55)

Follow-up of cohort members by calendar time and age

| entry

- exit because of disease onset (outcome of interest)
- exit due to other reason (censoring)

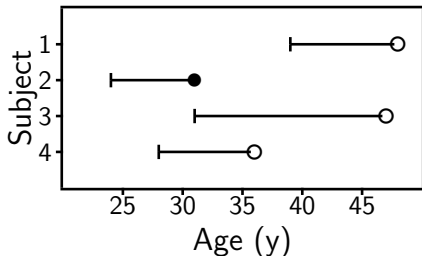
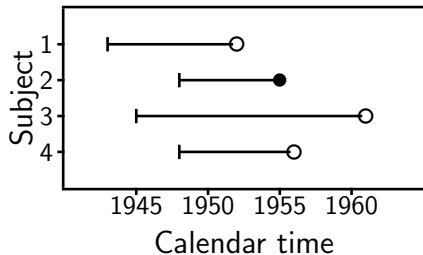


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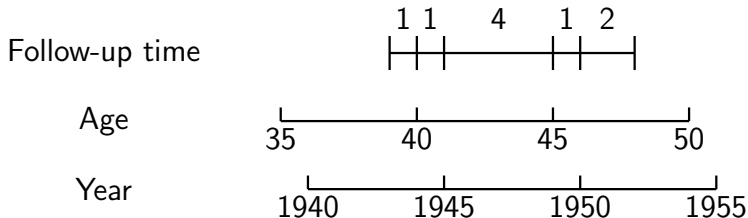
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Person-years by age and period

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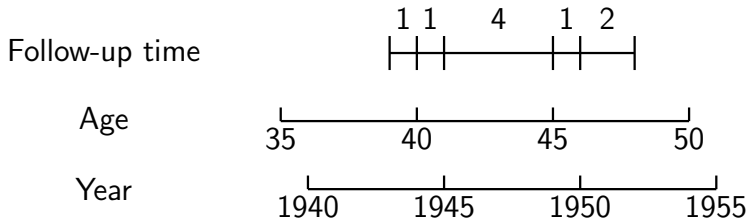
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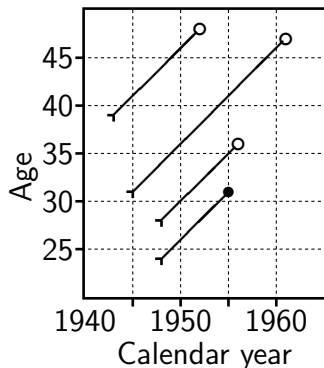
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This subject contributes person-time into 5 different cells defined by ageband & calendar period

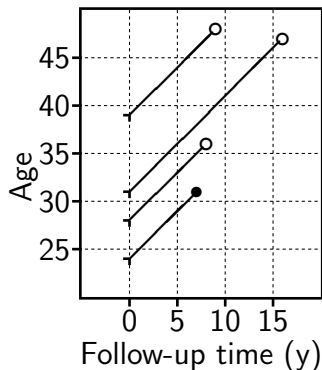
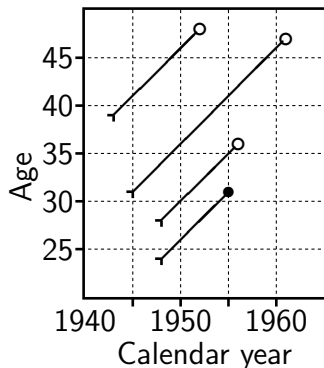
Follow-up in Lexis-diagrams

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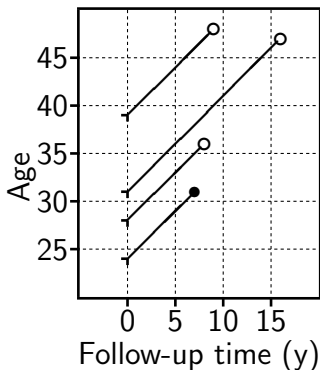
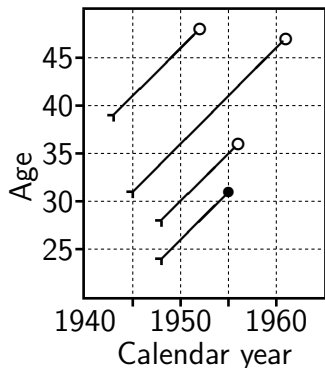
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Follow-up lines run diagonally through different ages and calendar periods.

See also Laufey's lecture on cohort studies, slide 4.

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- ▶ **Adjustment** or **standardization** for age needed!

Ex. Male stomach cancer in Cali and Birmingham (IS, Table 4.2, p. 71)

Age (y)	Cali			Birmingham			Rate ratio
	Male cases 1982 -86	Male Popu- lation 1984 ($\times 10^3$)	Incid. Rate ($/10^5y$) 1982 -86	Male cases 1983 -86	Male Popu- lation 1985 ($\times 10^3$)	Incid. Rate ($/10^5y$) 1983 -86	
39		524.2	1.5	79	1 683.6	1.2	<i>1.25</i>
266		76.3	69.7	1037	581.5	44.6	<i>1.56</i>
315		22.4	281.3	2352	291.1	202.0	<i>1.39</i>
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- ▶ In each age group Cali has a higher incidence but the crude incidence is higher in Birmingham.
- ▶ **Is there a paradox?**

Comparison of age structures

(IS, Tables 4.3,4.4)

Age (years)	% of male population			
	Cali 1984	B'ham 1985	Finland 1999	World Stand.
0-44	84	66	61	74
45-64	12	23	27	19
65+	4	11	12	7
All ages	100	100	100	100

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⇒ Any summary rate must be **adjusted for age**.

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- ▶ Standard population can be real (e.g. one of the populations under comparison, or their average) or fictitious (e.g. World Standard Population, WSP)

Some standard populations:

Age group (years)	African	World	European	Truncated
0	2 000	2 400	1 600	—
1–4	8 000	9 600	6 400	—
5–9	10 000	10 000	7 000	—
10–14	10 000	9 000	7 000	—
15–19	10 000	9 000	7 000	—
20–24	10 000	8 000	7 000	—
25–29	10 000	8 000	7 000	—
30–34	10 000	6 000	7 000	—
35–39	10 000	6 000	7 000	6 000
40–44	5 000	6 000	7 000	6 000
45–49	5 000	6 000	7 000	6 000
50–54	3 000	5 000	7 000	5 000
55–59	2 000	4 000	6 000	4 000
60–64	2 000	4 000	5 000	4 000
65–69	1 000	3 000	4 000	—
70–74	1 000	2 000	3 000	—
75–79	500	1 000	2 000	—
80–84	300	500	1 000	—
85+	200	500	1 000	—
Total	100 000	100 000	100 000	31 000

Stomach cancer in Cali & B'ham

Age-standardized rates by the World Standard Population:

Age	Cali		Birmingham	
	Rate ^a	Weight	Rate ^a	Weight
0-44	1.5 ×	0.74 = 1.11	1.2 ×	0.74 = 0.89
45-64	69.7 ×	0.19 = 13.24	44.6 ×	0.19 = 8.47
65+	281.3 ×	0.07 = 19.69	202.0 ×	0.07 = 14.14
Age-standardised rate		34.04		23.50

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- ▶ ASR in Cali higher – coherent with the age-specific rates.

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45-64	69.7 ×	0.19 = 13.24	44.6 ×	0.19 = 8.47
65+	281.3 ×	0.07 = 19.69	202.0 ×	0.07 = 14.14
Age-standardised rate		34.04	23.50	

- ▶ ASR in Cali higher – coherent with the age-specific rates.
- ▶ Summary rate ratio estimate: **standardized rate ratio**

$$\text{SRR} = 34.0/23.5 = 1.44$$

Stomach cancer in Cali & B'ham

Age-standardized rates by the World Standard Population:

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- ▶ Known as **comparative mortality figure (CMF)** when the outcome is death (from cause C or all causes).

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- ▶ Based on relation btw risk $F(t)$ and hazard $\lambda(t)$, or
Inc. prop'n = $1 - \exp(-\text{cum. rate}) \approx \text{cum. rate}$

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(IS, Fig 4.11, p. 77)

Age-group (years)	Incidence rate (per 100 000 pyrs)
0-4, . . . , 15-19	0.0
20-24, 25-29	0.1
30-34	0.9
35-39	3.5
40-44	6.7
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50-54	26.8
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60-64	87.2
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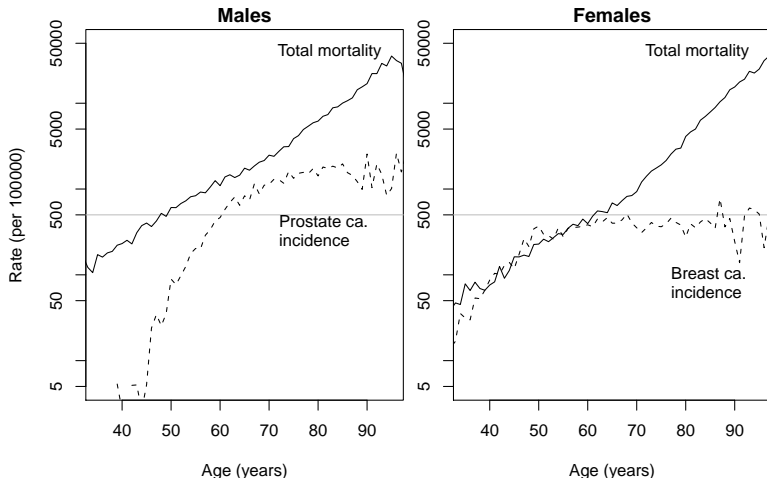
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- ▶ Age-specific and standardized rates are not very informative as such.
- ▶ Average cumulative risks are often estimated from cumulative rates.
- ▶ Yet, these estimates fictitiously presume that a person would not die from any cause before cancer hits him/her, but could even survive forever!

Total mortality and incidence of two common cancers by age, Finland 2005



Estimation of cumulative risks

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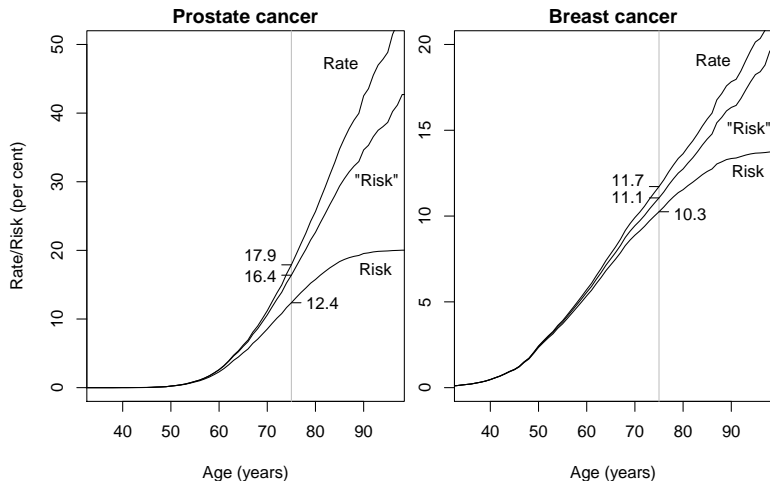
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- ▶ When this is properly done, the corrected estimates of cumulative risk will always be lower than the uncorrected “risks” .
- ▶ The magnitude of bias in the latter grows by age, but is reduced with increased life expectancy.

Cumulative measures, Finland 2005



Greater differences in males reflect shorter life expectancy and relatively high rates of prostate ca. in old ages.

Special cohorts of exposed subjects

- ▶ Occupational cohorts, exposed to potentially hazardous agents (e.g. rubber workers, see Laufey's lecture on cohort studies, slides 19-20)
- ▶ Cohorts of patients on chronic medication, which may have harmful long-term side-effects
- ▶ No internal comparison group of unexposed subjects.

Question: Do incidence or mortality rates in the *exposed* target cohort differ from those of a roughly comparable *reference* population?

Reference rates obtained from:

- ▶ population statistics (mortality rates)
- ▶ disease & hospital discharge registers (incidence)

Observed and expected cases – SIR

- ▶ Compare rates in a study cohort with a standard set of age-specific rates from the reference population.
- ▶ Reference rates normally based on large numbers of cases, so they are assumed to be “known” without error.
- ▶ Calculate **expected** number of cases, E , if the standard age-specific rates had applied in our study cohort.
- ▶ Compare this with the **observed** number of cases, D , by the **standardized incidence ratio** SIR
(or st'zed mortality ratio SMR with death as outcome)

$$\text{SIR} = D/E, \quad \text{SE}(\log[\text{SIR}]) = 1/\sqrt{D}$$

Example: HT and breast ca.

- ▶ A cohort of 974 women treated with hormone (replacement) therapy were followed up.
- ▶ $D = 15$ incident cases of breast cancer were observed.
- ▶ Person-years (Y) and reference rates (λ_a^* , per 100000 y) by age group (a) were:

Age	Y	λ_a^*	E
40–44	975	113	1.10
45–49	1079	162	1.75
50–54	2161	151	3.26
55–59	2793	183	5.11
60–64	3096	179	5.54
Σ			16.77

Ex: HT and breast ca. (cont'd)

- ▶ “Expected” cases at ages 40–44:

$$975 \times \frac{113}{100\,000} = 1.10$$

- ▶ Total “expected” cases is $E = 16.77$
- ▶ $SIR = 15/16.77 = 0.89$.
- ▶ Error-factor: $\exp(1.96 \times \sqrt{1/15}) = 1.66$
- ▶ 95% confidence interval is:

$$0.89 \overset{\times}{\div} 1.66 = (0.54, 1.48)$$

SIR for Cali with B'ham as reference

Total person-years at risk and expected number of cases in Cali 1982-86 based on age-specific rates in Birmingham
(IS: Fig. 4.9, p. 74)

Age	Person-years	Expected cases in Cali
0-44	$524\ 220 \times 5 = 2\ 621\ 100$	$0.000012 \times 2\ 621\ 100 = 31.45$
45-64	$76\ 304 \times 5 = 381\ 520$	$0.000446 \times 381\ 520 = 170.15$
65+	$22\ 398 \times 5 = 111\ 990$	$0.002020 \times 111\ 990 = 226.00$
	=3 114 610	Total expected (E) 427.82

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Standardised incidence ratio:

$$\text{SIR} = \frac{O}{E} = \frac{620}{427.8} = 1.45 \quad (\text{or } 145 \text{ per } 100)$$

Crude and adjusted rates compared

(IS: Table 4.6, p. 78, extended)

		Cali, 1982-86	B'ham, 1983-86	Rate ratio
Crude rates (/10 ⁵ y)		19.9	33.9	0.59
		48.0	33.9	1.42
ASR (/10 ⁵ y) ^C	—"–	19.9	14.4	1.38
ASR (/10 ⁵ y) ^W	—"–	34.0	23.5	1.44
		14.6	9.5	1.54
		36.3	21.2	1.71
Cum. rate < 75 y (per 1000)	—"–	46.0	26.0	1.77

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NB: The ratios of age-adjusted rates appear less dependent on the choice of standard weights than on the coarseness of age grouping. 5-year age groups are preferred.

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- ▶ outcome event of interest = death,
- ▶ measures and methods used somewhat different from those used in incidence analysis.

Follow-up of 8 out of 40 breast cancer patients (from IS, table 12.1., p. 264)

No.	Age (y)	Stage ^a	Date of diagnosis	Date at end of follow-up	Vital status at end of follow-up	Cause of death ^c	Full years from diagn's up to end of follow-up	Days from diagn's up to end of follow-up
	39	1	01/02/89	23/10/92	A	–	3	1360
3	56	2	16/04/89	05/09/89	D	BC	0	142
5	62	2	12/06/89	28/12/95	A	–	6	2390
15	60	2	03/08/90	27/11/94	A	–	4	1577
22	64	2	17/02/91	06/09/94	D	O	3	1297
25	42	2	20/06/91	15/03/92	D	BC	0	269
30	77	1	05/05/92	10/05/95	A	–	3	1100
37	45	1	11/05/93	07/02/94	D	BC	0	272

^a 1 = absence of regional lymph node involvement and metastases

2 = involvement of regional lymph node and/or presence of metastases

^b A = alive; D = dead; ^c BC = breast cancer; O = other causes

Follow-up of 8 out of 40 breast cancer patients (from IS, table 12.1., p. 264)

No.	Age (y)	Stage ^a	Date of diagnosis	Date at end of follow-up	Vital status at end of follow-up	Cause of death ^c	Full years from diagn's up to end of follow-up	Days from diagn's up to end of follow-up
1	39	1	01/02/89	23/10/92	A	–	3	1360
3	56	2	16/04/89	05/09/89	D	BC	0	142
5	62	2	12/06/89	28/12/95	A	–	6	2390
15	60	2	03/08/90	27/11/94	A	–	4	1577
22	64	2	17/02/91	06/09/94	D	O	3	1297
25	42	2	20/06/91	15/03/92	D	BC	0	269
30	77	1	05/05/92	10/05/95	A	–	3	1100
37	45	1	11/05/93	07/02/94	D	BC	0	272

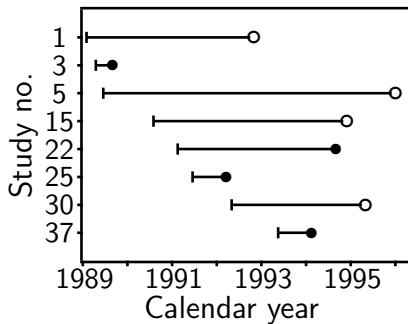
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Follow-up of breast ca. pts (cont'd)

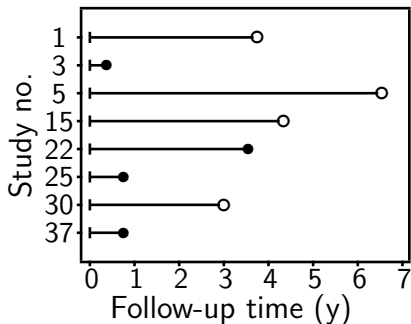
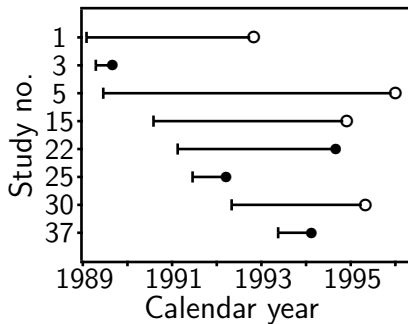
| entry = diagnosis; ● exit = death; ○ exit = censoring



(IS: Figure 12.1, p. 265)

Follow-up of breast ca. pts (cont'd)

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(IS: Figure 12.1, p. 265)

Life table or “actuarial” method

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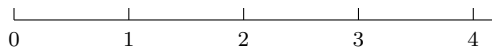
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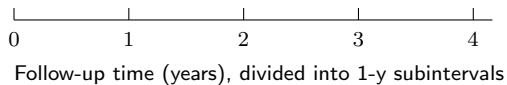
$L_k =$ no. of **losses**, *i.e.* individuals **censored** during the interval before being observed to die.

Life table items in a tree diagram

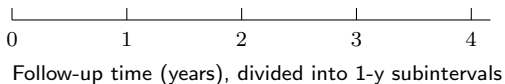
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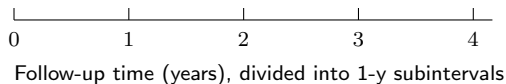


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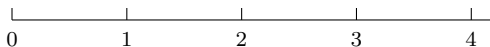


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at risk \square
 N_1

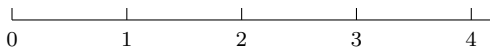
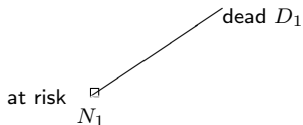


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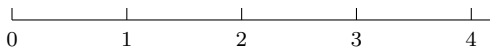
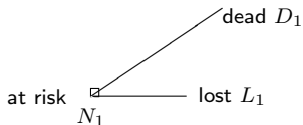


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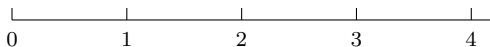
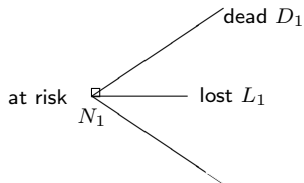


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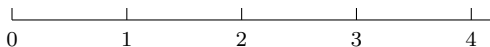
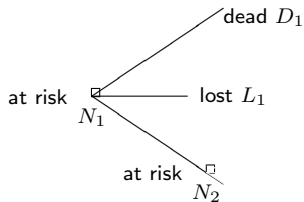


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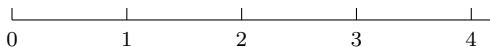
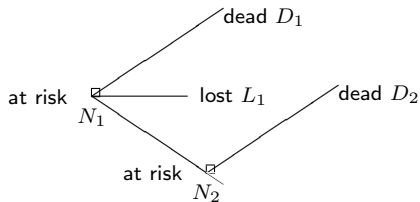


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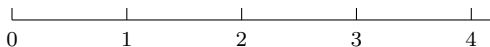
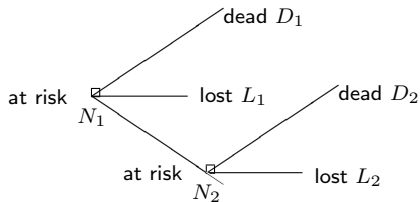


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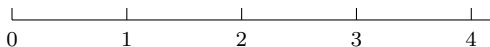
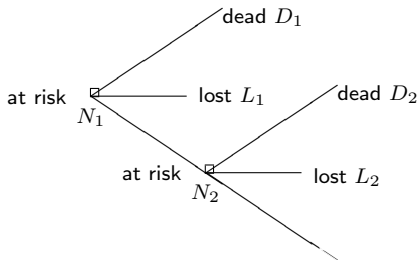


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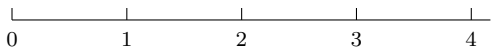
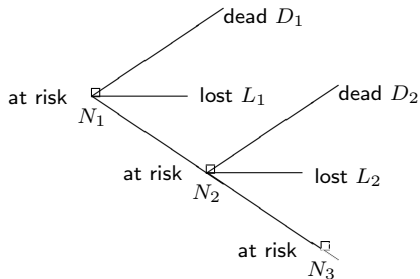


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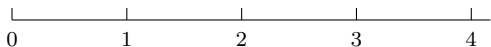
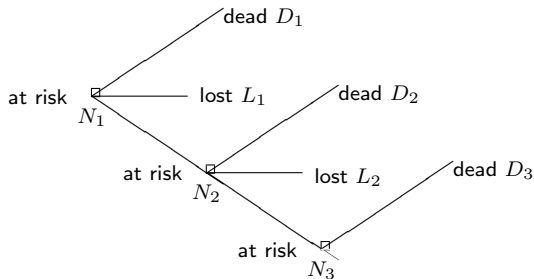


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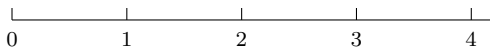
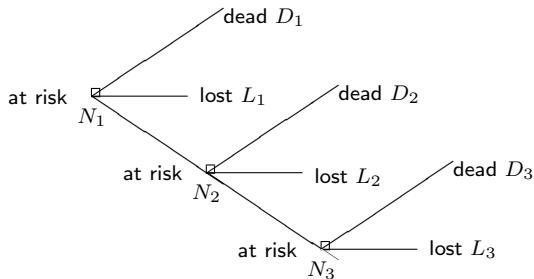


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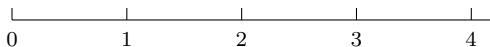
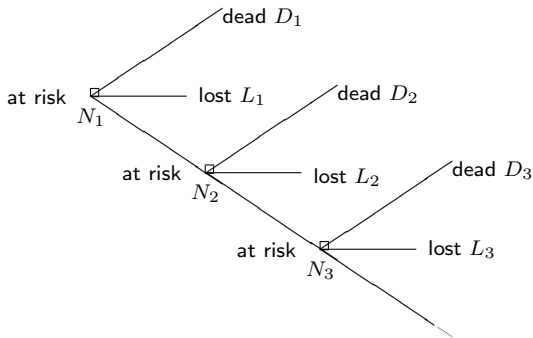


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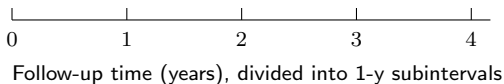
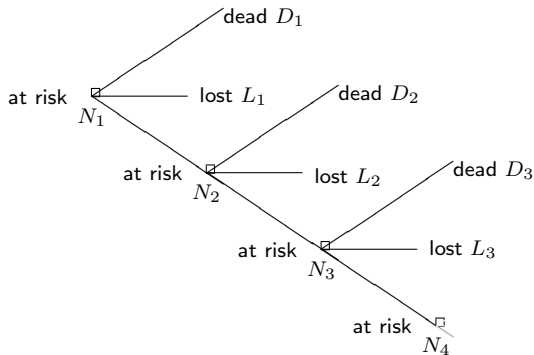


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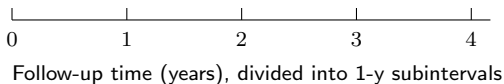
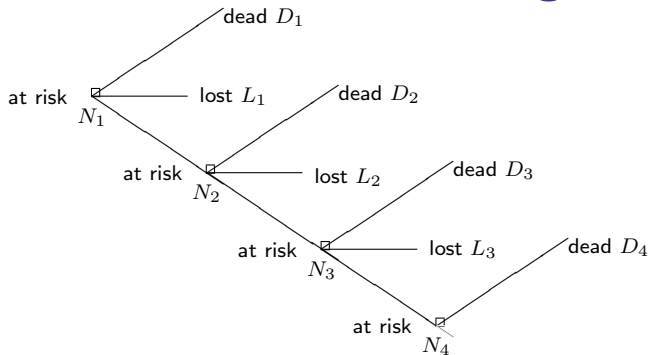
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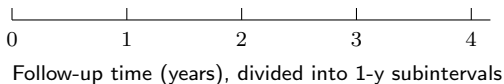
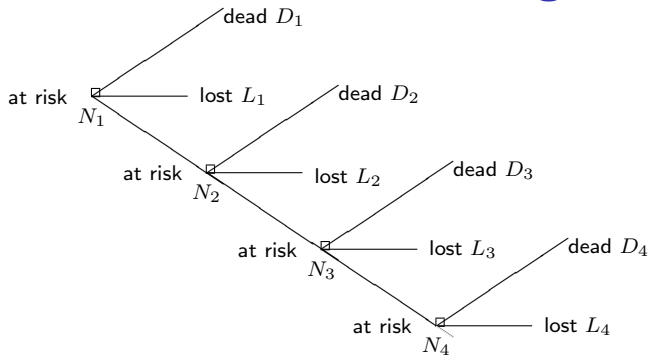
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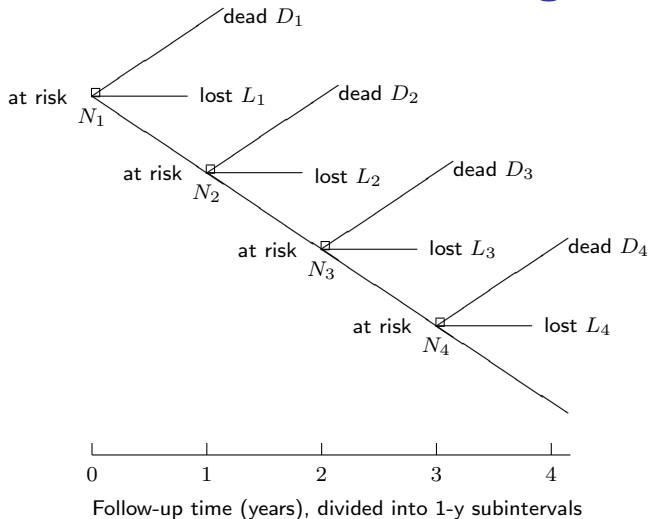
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Life table items for breast ca. patients

(IS: Table 12.2., p. 273, first 4 columns)

Inter- val (k)	Years since diagnosis	No. at start of interval (N_k)	No. of deaths (D_k)	No. of losses (L_k)
1	0- < 1	40	7	0
2	1- < 2	33	3	6
3	2- < 3	24	4	3
4	3- < 4	17	4	4
5	4- < 5	9	2	3
6	5- < 6	4	1	2
7	6- < 7	1	0	1
Total			21	19

Life table calculations (cont'd)

(3) Calculate and tabulate for each interval

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date of diagnosis until the end of the k th interval

$=$ estimate of **survival probability** up to this time point.

Follow-up of breast ca. patients (cont'd)

Actuarial life table completed (IS, table 12.2, p. 273)

Inter- val	Years since dia- gnosis	No. at start of in- terval	No. of deaths	No. of losses	Effec- tive deno- minator	Cond'l prop'n of deaths during int'l	Survival prop'n over int'l	Cumul. survival; est'd survival prob'ty
(k)		(N_k)	(D_k)	(L_k)	(N'_k)	(q_k)	(p_k)	(S_k)
1	0- < 1	40	7	0	40.0	0.175	0.825	0.825
2	1- < 2	33	3	6	30.0	0.100	0.900	0.743
3	2- < 3	24	4	3	22.5	0.178	0.822	0.610
4	3- < 4	17	4	4	15.0	0.267	0.733	0.447
5	4- < 5	9	2	3	7.5	0.267	0.733	0.328
6	5- < 6	4	1	2	3.0	0.333	0.667	0.219
7	6- < 7	1	0	1	0.5	0.0	1.0	0.219

Follow-up of breast ca. patients (cont'd)

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1-year survival probability is thus estimated 82.5% and 5-year probability 32.8%.

Comparison to previous methods

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- ▶ “Actuarial” incidence rate in the k th interval:

$$I_k = \frac{\text{number of cases } (D_k)}{\text{approximate person-time}}$$

where the person-time is approximated by

$$\left[N_k - \frac{1}{2}(D_k + L_k) \right] \times \text{length of interval}$$

The dead and censored thus contribute half of the interval length.

Survival curve and other measures

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Line diagram of survival proportions through interval endpoints

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Line diagram of survival proportions through interval endpoints provides graphical estimates of interesting parameters of the survival time distribution, e.g.:

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- ▶ **median** and **quartiles**: time points at which the curve crosses the 50%, 75%, and 25% levels
- ▶ **mean residual lifetime**: area under the curve, given that it decreases all the way down to the 0% level.

NB. Often the curve ends at higher level than 0%, in which case some measures cannot be calculated.

Survival curve of breast ca. patients (IS: Fig 12.8)

Numbers above x -axis show the size of population at risk.

Cause-specific and relative survival

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(B) **Relative survival** analysis: Compute

$$R_k = S_k^{\text{obs}} / S_k^{\text{exp}},$$

the **relative survival proportion** = ratio of

- ▶ **observed** survival proportion S_k^{obs} in cancer patients,
- ▶ **expected** survival proportion S_k^{exp} based on age-specific mortalities in a reference population (*cf.* SIR!)

Cause-specific and relative survival

(A) **Cause-specific** survival analysis:

- ▶ outcome event: death from the disease C itself,
- ▶ deaths from other causes \rightarrow counted as losses,
- problems with cause of death & competing causes.

(B) **Relative survival** analysis: Compute

$$R_k = S_k^{\text{obs}} / S_k^{\text{exp}},$$

the **relative survival proportion** = ratio of

- ▶ **observed** survival proportion S_k^{obs} in cancer patients,
 - ▶ **expected** survival proportion S_k^{exp} based on age-specific mortalities in a reference population (*cf.* SIR!)
- + no information on causes of death needed.

Ex. Breast cancer patients (cont'd)

Overall and cause-specific (death from breast ca.) survival
(**IS**: Fig 12.9 & 12.12, p. 271-3)

Kaplan-Meier curves – alternative to “actuarial”:

NB. Meaning of “cause-specific survival”?

CONCLUSION

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Major challenges:

- ▶ obtain the right denominator for each numerator,
- ▶ valid calculation of person-years,
- ▶ appropriate treatment of time and its various aspects,
- ▶ removal of confounding from comparisons.